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# MODIFIED ROBUST RATIO-REGRESSION ESTIMATORS FOR POPULATION MEAN WITH AUXILIARY VARIABLE

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#### **ABSTRACT**

In sample survey, the uses of auxiliary variable in estimating population mean has been paramount for many researchers, because, it improves the efficiency of proposed estimators significantly. Several studies show that higher level of precision could be obtained when the auxiliary information were adopted especially when both variables were highly correlated. In this study, regression approach and sample size of auxiliary variable were used to improve the precision of proposed estimators. The bias and mean square error of each proposed modified ratio estimators using regression approach were obtained up to first order approximation. A theoretical comparisons of propose modified ratio estimators were carried out with other existing estimators for population mean. A numerical study was also carried out to see the performance and efficiency of the proposed estimators over some other existing estimators using Murthy (1967) and Mokhopadhyan (2009) data set. And the condition under which the proposed modified ratio estimators with the regression approach better than other existing modified ratio estimators using regression approach were better than other existing modified ratio estimators has lower mean square error compared to other existing estimators.

**Keywords**: Ratio estimator, Regression Approach, Sample Size, Bias, Mean Square Error, Efficiency.

#### 1. Introduction

The use of auxiliary information to increase or improve the efficiency of suggested or modified estimator in sampling theory has been a custom to many researchers in sample survey. In fact it more efficient when there are higher level of positive correlation between a studying variable and auxiliary. Auxiliary information incorporated with study variable contribute immensely in the estimation of population mean in finite population if its use properly. However, research exhibits that the uses of auxiliary variable incorporated with the study variable to modify ratio type estimator perform efficiently even when there existing negative correlation between the two variables. Many researchers have proposed an transformed estimators by conventional estimators such as ratio, product, different and regression estimators to relative type with adoption of auxiliary information in order to increase the precision of the estimate of the population mean.

The sample mean is the most acceptable and suitable estimator to estimate population mean when the sample size is large but research shows that regression approach play significance roles in efficiency of any estimator. Cochran W. G (1940) and Cochran WG (1977) initiated the

use of auxiliary information and proposed a ratio estimator for population mean. It is so far established fact that the ratio type estimator provides better efficiency in comparison to simple mean estimator provided the study and auxiliary variables are positively correlated. If the correlation between the study and auxiliary variables is negative, product estimator is more efficient than sample mean estimator. Modified ratio estimators came into existence and were constructed by using one or more unknown constants. In a class of estimators, the estimator with minimum variance or mean square error is regarded as the most efficient estimator. This concept has been utilized by several researchers to improve the efficiency of ratio and product type estimators for estimating population mean of study variable like Upadhyaya L.N and Singh HP (1999), Singh H. O., Tailor R., Kakran M. S. (2004), Jeelani M. I., Magbool S., Mir S.A (2013), Yadav S. K., Shukla A.K (2014), Abid M, Abbas N, Sherwani RAK and Nazir HZ (2016), Yadav S.K and Pandey H(2011), ,Yadav S.K and AdewaraA.A (2013), Yadav S.K, Mishra S.S and Shula A.K (2014), Yadav S.K, Mishra S.S and Shula A.K (2015), Yadav S.K, Mishra S.S, Shula A.K, Kumar S and Singh R.S (2016a), Yadav S.K, Gupta S.A.T, Mishra S.S, and Shula A.K





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(2016b), Yadav S.K, Subramani J, Mishra S.S, and Shula A.K (2016c), Yadav S.K, Misra S, Mishra S.S, and Chutiman(2016d)

Let  $U=\{U_1,\ U_2,\ U_3,...,U_N\}$  be a finite population having units and each  $U_i=(X_i,Y_i)$ , i=1,2,3,...,N has a pair of values. Y is the study variable and X is the auxiliary variable which is correlated with Y. Let  $y=\{y_1,y_2,y_3,...,y_n\}$  and  $x=\{x_1,x_2,x_3,...,x_n\}$  be n sample values,  $\overline{y}$  and  $\overline{x}$  are the sample means of the study and auxiliary variables respectively. Let  $S_Y^2$  and  $S_X^2$  be the finite population variance of Y and X respectively and  $S_y^2$  and  $S_x^2$  be respective sample variances based on the random sample of size n drawn without

N: Population size, n: Sample size,

Y: Study variable, X: Auxiliary variable

 $\bar{y}, \bar{x}$ : Sample means of study and auxiliary variables.

 $\bar{Y}, \bar{X}$ : Population means of study and auxiliary variables.

ρ:Coefficient of correlation of study and auxiliary variables

 $C_y$ ,  $C_x$ : Coefficient of variations of study and auxiliary variables

 $\beta_1$ : Coefficient of skweness of auxiliary variable.  $\beta_2$ : Coefficient of kurtosis of auxiliary variable,  $M_d$ : Median of the auxiliary variable

$$S_{y}^{2} = \frac{1}{n-1} \sum_{i=0}^{n} (y_{i} - \bar{y})^{2}, \quad S_{x}^{2} = \frac{1}{n-1} \sum_{i=0}^{n} (x_{i} - \bar{x})^{2}, \quad S_{Y}^{2} = \frac{1}{N-1} \sum_{i=0}^{N} (Y_{i} - \bar{Y})^{2},$$

$$S_{X}^{2} = \frac{1}{N-1} \sum_{i=0}^{n} (X_{i} - \bar{X})^{2}, \quad \theta = \left(\frac{1}{n} - \frac{1}{N}\right), \quad C_{y} = \frac{S_{y}}{\bar{Y}^{2}}, \quad C_{x} = \frac{S_{x}}{\bar{X}^{2}}.$$

### $S_X - \frac{1}{N-1} \sum_{i=0}^{N-1} (A_i - A_i), \quad 0 - \left(\frac{1}{n} - \frac{1}{N}\right),$

### 1. Literature review

The most suitable estimator for estimating population mean  $\bar{Y}$  is the sample  $\bar{y}$  given by

$$\bar{y} = \frac{1}{n} \sum_{i=0}^{n} y_i \tag{1}$$

And it is unbiased for population mean and its variance up to first order of approximation this is given by,  $V(\bar{y}) = \theta S_v^2 = \theta \bar{Y}^2 C_v^2$ , (2)

Cochran (1940) proposed an estimator by using positive correlation of auxiliary variable with study variable and the following usual ratio estimator of population mean as,

$$\widehat{\overline{Y}} = \overline{y}.\frac{\overline{X}}{\overline{Y}}$$
 (3)

The above estimator is biased estimator of population mean and its Bias and Mean Square Error were obtained up to first order approximation respectively.

Bias 
$$(\hat{\mathbf{Y}}_r) = \theta \bar{Y} \left( C_x^2 - \rho C_x C_y \right)$$
 (4)

$$MSE (\hat{\mathbf{Y}}_{r}) = \theta \bar{Y}^{2} \left( C_{x}^{2} + C_{y}^{2} - 2\rho C_{x} C_{y} \right)$$

$$\tag{5}$$

Sisodia B. V. S. and Dwivedi V.K (1981) proposed ratio type estimator by modified auxiliary information with the uses of coefficient of variation ( $C_x$ ) of auxiliary variable as:

$$\widehat{\bar{Y}}_1 = \bar{y} \left( \frac{\bar{X} + C_X}{\bar{X} + C_X} \right) \tag{6}$$

Bias 
$$(\hat{Y}_1) = \theta \bar{Y} \left( \delta_1^2 C_x^2 - 2\rho C_x C_y \right)$$
 (7)

$$MSE\left(\widehat{Y}_{1}\right) = \theta \overline{Y}^{2} \left(C_{y}^{2} + \delta_{1}^{2} C_{x}^{2} + -2\rho C_{x} C_{y}\right), \tag{8}$$

where it's constant is given as: 
$$\delta_1 = \frac{\bar{\chi}}{\bar{\chi} + C_{\chi}}$$
 (9)

Upadhyaya L. N., Singlh H. P,(1999) proposed ratio type estimator by imposing coefficient of variation,  $C_x$  on work of Singh H. P., Tailor R., and Kakran M. S. (2004), the propose estimator it as:

$$\widehat{\bar{Y}}_5 = \bar{y} \left( \frac{\bar{X}C_x + \beta_2}{\bar{x} C_x + \beta_2} \right) \tag{10}$$

Bias 
$$(\widehat{Y}_5) = \theta \overline{Y} \left( \delta_5^2 C_x^2 - 2\rho C_x C_y \right)$$
 (11)

$$MSE\left(\widehat{Y}_{5}\right) = \theta \overline{Y}^{2} \left(C_{y}^{2} + \delta_{5}^{2} C_{x}^{2} + -2\rho C_{x} C_{y}\right)$$

$$\tag{12}$$

where it's constant is given as: 
$$\delta_5 = \frac{\bar{X}C_X}{\bar{X} + \beta_2}$$
 (13)

Singh H. P., Tailor R. (2003) proposed ratio type estimator by replacing correlation of an auxiliary variable on work of Sisodia B. V. S. and Dwivedi V.K (1981)

$$\widehat{\bar{Y}}_2 = \bar{y} \left( \frac{\bar{X} + \rho}{\bar{x} + \rho} \right) \tag{14}$$

Bias 
$$(\hat{\bar{Y}}_2) = \theta \bar{Y} \left( \delta_2^2 C_x^2 - 2\rho C_x C_y \right)$$
 (15)





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$$MSE\left(\widehat{Y}_{2}\right) = \theta \overline{Y}^{2} \left(C_{y}^{2} + \delta_{2}^{2} C_{x}^{2} + -2\rho C_{x} C_{y}\right), \tag{16}$$

where it's constant is given as: 
$$\delta_2 = \frac{\bar{X}}{\bar{X} + \rho}$$
 (17)

Singh H. P., Tailor R., and Kakran M. S. (2004) proposed ratio type estimator by replacing coefficient of variation ( $C_x$ ) in work of Sisodia B. V. S. and Dwivedi V.K (1981) with coefficient of kurtosis ( $\beta_2$ ) of auxiliary variable as:

$$\widehat{\bar{Y}}_6 = \bar{y} \left( \frac{\bar{X} + \beta_2}{\bar{X} + \beta_2} \right) \tag{18}$$

Bias 
$$(\widehat{\overline{Y}}_6) = \theta \overline{Y} \left( \delta_6^2 C_x^2 - 2\rho C_x C_y \right)$$
 (19)

MSE 
$$(\hat{\bar{Y}}_6) = \theta \bar{Y}^2 (C_y^2 + \delta_6^2 C_x^2 + -2\rho C_x C_y)$$
 (20)

where it's constant is given as: 
$$\delta_6 = \frac{\bar{X}}{\bar{X} + \beta_2}$$
 (21)

Subramani J., Kumarpandiyan G.(2013) proposed an estimator by replacing the correlation (ρ) on the work of Singh H. P., Tailor R. (2003), and coefficient of kurtosis ( $\beta_2$ ) on the work of Upadhyaya L. N., Singlh H. P, (1999) by median of auxiliary variable.

$$\widehat{\bar{Y}}_{3} = \bar{y} \left( \frac{\bar{X} + M_d}{\bar{x} + M_d} \right) \tag{22}$$

Bias 
$$(\hat{\bar{Y}}_3) = \theta \bar{Y} \left( \delta_3^2 C_x^2 - 2\rho C_x C_y \right)$$
 (23)

$$MSE\left(\widehat{Y}_{3}\right) = \theta \overline{Y}^{2} \left(C_{y}^{2} + \delta_{3}^{2} C_{x}^{2} + -2\rho C_{x} C_{y}\right), \tag{24}$$

MSE 
$$(\hat{Y}_3) = \theta \bar{Y}^2 (C_y^2 + \delta_3^2 C_x^2 + -2\rho C_x C_y)$$
, where it's constant is given as:  $\delta_3 = \frac{\bar{X}}{\bar{X} + M_d}$  (25)

Jerajuddin and Kishun (2016) proposed a ratio type estimator that minimized an error term and biasness compare to other ratio type. They are the first to use the sample size n in order to increase the precision of Cochran W.G (1940) estimator. The estimator it as:

$$\widehat{\overline{Y}}_4 = \bar{y} \left( \frac{\bar{X}C_X + n}{\bar{x}C_X + n} \right) \tag{26}$$

Bias 
$$(\widehat{Y}_4) = \theta \, \overline{Y} \left( \delta_{JKP}^2 C_x^2 - 2\rho C_x C_y \right)$$
 (27)

$$MSE\left(\hat{Y}_{4}\right) = \theta \bar{Y}^{2} \left(C_{y}^{2} + \delta_{JKP}^{2} C_{x}^{2} - 2\rho C_{x} C_{y}\right)$$

$$(28)$$

where it's constant is given as: 
$$\delta_4 = \frac{\bar{X}C_X}{\bar{X} + \beta_2}$$
 (29)

### 2. Researched Methodology

Suleiman and Adewara (2021), proposed an improved modified ratio estimator of population mean using information on size of the sample. The bias and mean square error up to first order

$$\zeta_{pi} = \overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\overline{x} + C_{xn}}{\overline{x} + C_{xn}} \right) \right]$$

of approximation were obtained. The generalized estimators of the population mean, and it's constant, bias and mean square error where is given as follows:

(30)

Table 1: Biases and Mean Square Errors (MSE) of the existing estimators

	Estimator(s) $\zeta_{pi}$	Constant	Bias	MSE
s/n	•	$\delta_{pi}$		
1	$\overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\bar{X} + C_{xn}}{\bar{x} + C_{xn}} \right) \right]$	$\frac{\bar{X}}{\bar{X} + C_{x n}}$	$\theta  \overline{y} \left[ -\delta_{p1} \rho C_y C_x + \delta_{p1}^2 C_x^2 + \alpha_1 \delta_{p1} \rho C_y C_x - \alpha_1 \delta_{p1}^2 C_x^2 \right]$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + \delta_{p1}^{2}C_{x}^{2} - 2\delta_{p1}\rho C_{y}C_{x} - \frac{A_{1}^{2}}{B_{1}^{2}} \right]$
2	$ \overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\overline{X}C_x + \beta_2 n}{\overline{x}C_x + \beta_2 n} \right) \right] $	$\frac{\bar{X}C_x}{\bar{X}C_x + \beta_2 n}$	$\theta \overline{Y} \left[ -\delta_{p2} \rho C_y C_x + \delta_{p2}^2 C_x^2 + \alpha_2 \delta_{pi} \rho C_y C_x - \alpha_2 \delta_{p2}^2 C_x^2 \right]$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + \delta_{p2}^{2}C_{x}^{2} - 2\delta_{p2}\rho C_{y}C_{x} - \frac{A_{2}^{2}}{B_{2}^{2}} \right]$



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3	$ \overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\overline{X} + \rho n}{\overline{x} + \rho n} \right) \right] $	$\frac{\bar{X}}{\bar{X} + \rho n}$	$\theta \bar{Y} \left[ -\delta_{p3}\rho C_y C_x + \delta_{p3}^2 C_x^2 + \alpha_3 \delta_{p3} \rho C_y C_x - \alpha_3 \delta_{p3}^2 C_x^2 \right]$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + \delta_{p3}^{2}C_{x}^{2} - 2\delta_{p3}\rho C_{y}C_{x} - \frac{A_{3}^{2}}{B_{3}^{2}} \right]$
4	$\overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\overline{X} + \beta_2 n}{\overline{x} + \beta_2 n} \right) \right]$	$\frac{\bar{X}}{\bar{X} + \beta_2 n}$	$\theta \bar{Y} \left[ -\delta_{p4} \rho C_y C_x + \delta_{p4}^2 C_x^2 + \alpha_4 \delta_{p4} \rho C_y C_x - \alpha_4 \delta_p^2 C_x^2 \right]$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + \delta_{p4}^{2}C_{x}^{2} - 2\delta_{p4}\rho C_{y}C_{x} - \frac{A_{4}^{2}}{B_{4}^{2}} \right]$
5	$\overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\overline{X} + \beta_1 n}{\overline{x} + \beta_1 n} \right) \right]$	$\frac{\bar{X}}{\bar{X} + \beta_1 n}$	$\theta \bar{Y} \left[ -\delta_{p5}\rho C_y C_x + \delta_{p5}^2 C_x^2 + \alpha_5 \delta_{p5} \rho C_y C_x - \alpha_5 \delta_{p5}^2 C_x^2 \right]$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + \delta_{p5}^{2}C_{x}^{2} - 2\delta_{p5}\rho C_{y}C_{x} - \frac{A_{5}^{2}}{B_{5}^{2}} \right]$
6	$\overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\overline{X} + M_d n}{\overline{x} + M_d n} \right) \right]$	$\frac{\bar{X}}{\bar{X} + M_d n}$	$\theta \bar{Y} \left[ -\delta_{p6} \rho C_y C_x + \delta_{p6}^2 C_x^2 + \alpha_6 \delta_{p6} \rho C_y C_x - \alpha_6 \delta_{p6}^2 C_x^2 \right]$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + \delta_{p6}^{2}C_{x}^{2} - 2\delta_{p6}\rho C_{y}C_{x} - \frac{A_{6}^{2}}{B_{6}^{2}} \right]$
7	$ \overline{y} \left[ \alpha_i + (1 - \alpha_i) \left( \frac{\overline{X} M_d + n}{\overline{X} M_d + n} \right) \right] $	$\frac{\bar{X}}{\bar{X}M_dn+n}$	$\theta \overline{Y} \left[ -\delta_{p7} \rho C_y C_x + \delta_{p7}^2 C_x^2 + \alpha_7 \delta_{p7} \rho C_y C_x - \alpha_7 \delta_{p7}^2 C_x^2 \right]$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + \delta_{p7}^{2}C_{x}^{2} - 2\delta_{p7}\rho C_{y}C_{x} - \frac{A_{7}^{2}}{B_{7}^{2}} \right]$

### 3. The proposed estimators;

This research work was motivated by the work of Suleiman and Adewara (2021), and propose

improved modified ratio estimation of population mean using regression approach. The estimators are;

$$K_1 = \overline{y}[\alpha_1 + b\varphi(\overline{X} - \overline{x})] \exp\left[\frac{\overline{X} + C_{Xn}}{\overline{x} + C_{Xn}}\right]$$
(31)

$$K_{2} = \overline{y}[\alpha_{2} + b\varphi (\overline{X} - \overline{x})] \exp \left[\frac{\overline{X}C_{X} + \beta_{2}n}{\overline{x}C_{X} + \beta_{2}n}\right]$$
(32)

$$K_3 = \overline{y}[\alpha_3 + b\varphi(\overline{X} - \overline{x})] \exp\left[\frac{\overline{X} + \rho n}{\overline{x} + \rho n}\right]$$
(33)

$$K_4 = \overline{y}[\alpha_4 + b\varphi(\overline{X} - \overline{x})] \exp\left[\frac{\overline{X} + \beta_2 n}{\overline{x} + \beta_2 n}\right]$$
(34)

$$K_{5} = \overline{y}[\alpha_{5} + b\varphi(\overline{X} - \overline{x})] \exp\left[\frac{\overline{X} + \beta_{1}n}{\overline{x} + \beta_{1}n}\right]$$
(35)

$$K_6 = \overline{y}[\alpha_6 + b\varphi(\overline{X} - \overline{x})] exp\left[\frac{\overline{X} + M_d n}{\overline{x} + M_d n}\right]$$
(36)

$$K_7 = \overline{y}[\alpha_7 + b\varphi(\bar{X} - \bar{x})] \exp\left[\frac{\bar{X}M_d + n}{\bar{x}M_d + n}\right]$$
(37)

Where,  $\alpha_i (i=1,2,3,4,5,6,7)$  are suitably chosen constant to be defined such that the mean square error of the proposed estimator is minimum and  $b_{\varphi}$  is slope of the regression curve, this could be obtained by regression of both study and auxiliary variable.

To obtain the bias and Mean Square Error of the propose estimators, the following properties must be adopted.

Properties of the Proposed Estimators

$$\bar{y} = \bar{Y} (1 + e_0), \ \bar{x} = \bar{X} (1 + e_1) \text{ such that } E(e_0) = E(e_1) = 0, \ E(e_0^2) = \theta \ C_y^2, \ E(e_1^2) = \theta C_x^2, \ E(e_0 e_1) = \theta \rho C_y C_x, \text{ where } \theta = \frac{1-f}{n}$$
 (38)

Bias and MSE of  $K_1$ 





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Now express eq (21) in error term and substitute eq (28) into equation eq (31)

$$K_{1} = \bar{Y}(1 + e_{0}) \left[ \alpha + b\varphi \left( \bar{X} - \bar{X} (1 + e_{1}) \right) \exp \left[ \frac{\bar{X} + C_{X} n}{\bar{X} (1 + e_{1}) + C_{X} n} \right]$$
(39)

$$K_1 = \overline{Y}(1 + e_0) \left[\alpha + b\varphi \left(\overline{X} - \overline{X} - \overline{X}e_1\right)\right] \exp \left[\frac{\overline{X} + C_{XR}}{\overline{X} + \overline{X}e_1 + C_{XR}}\right]$$

$$\tag{40}$$

$$K_1 = \overline{Y}(1 + e_0) \left[ \alpha - b\varphi \overline{X} e_1 \right] \exp \left[ \frac{\overline{X} + C_{Xn}}{\overline{X} + C_{Xn} + \overline{X} e_1} \right]$$

$$\tag{41}$$

$$K_1 = \overline{Y}(1 + e_0) \left[\alpha - b\varphi \overline{X}e_1\right] \exp\left[\frac{1}{1 + \frac{\overline{X}}{\overline{X} + C_X n}e_1}\right]$$
(42)

Use inverse operation for eq (42), we have

$$K_{1} = \overline{Y} + \overline{Y}e_{0}\left[\alpha + b\varphi\left(\overline{X} - \overline{X}\left(1 + e_{1}\right)\right]\exp\left[\frac{1}{1 + \frac{\overline{X}}{\overline{X} + C_{X}n}e_{1}}\right]$$

$$\tag{43}$$

$$K_1 = \overline{Y}(1 + e_0) \left[\alpha - b\varphi \overline{X}e_1\right] \exp\left[\frac{1}{1 + \delta_p e_1}\right]$$
(44)

$$K_1 = \overline{Y}(1 + e_0) \left[\alpha - b\varphi \overline{X}e_1\right] \exp\left[\left(1 + \delta_p e_1\right)^{-1}\right]$$
(45)

Where it constant is: 
$$\delta_p = \frac{\bar{X}}{\bar{X} + C_{X,n}}$$
 (46)

$$K_{1} = \overline{Y}(1 + e_{0}) \left[\alpha - b\varphi \overline{X}e_{1}\right] \cdot \exp\left[\left(1 + \delta_{p}e_{1} + \delta_{p}^{2}e_{1}^{2}\right)^{-1}\right]$$
(47)

Retaining the term up to first order approximation, we have

$$K_{1} = (\bar{Y} + \bar{Y}e_{0})[\alpha + \alpha b\varphi \bar{X}e_{1} + \alpha \delta_{p}^{2}e_{1}^{2} - b\varphi \bar{X}e_{1} - b\varphi \bar{X}e_{1}\delta_{p}e_{1} - b\varphi \bar{X}e_{1}\delta_{p}^{2}e_{1}^{2}]$$
(48)

Simplify the equation and subtract  $\bar{Y}$  from both side of equation, we get

Simplify the equation and subtract 
$$\bar{Y}$$
 from both side of equation, we get 
$$K_1 - \bar{Y} = \bar{Y} \begin{bmatrix} \alpha + \alpha b \varphi \bar{X} e_1 + \alpha \delta_p^2 e_1^2 - b \varphi \bar{X} e_1 - b \varphi \bar{X} e_1 \delta_p e_1 - b \varphi \bar{X} e_1 \delta_p^2 e_1^2 + \alpha e_0 + \alpha b \varphi \bar{X} e_1 e_0 + \beta \bar{X} e_1 e_0 + \beta \bar{X} e_1 e_0 - \beta \bar{X} e_1 e_0 - \beta \bar{X} e_1 e_0 e_1 - \beta \bar{X} e_1 e_0 e_1 - \beta \bar{X} e_1 e_0 e_1 e_0 e_1 - \beta \bar{X} e_1 e_0 e_1 e_$$

Taking the expectation of both sides of equation and substitute the values of different expectations, we

Taking the expectation of both sides of equation and substitute the values of different expectations, we have bias of 
$$K_1$$

$$E(K_1 - \bar{Y}) = \bar{Y}E\begin{bmatrix} \alpha + \alpha b \phi \bar{X} e_1 + \alpha \delta_p^2 e_1^2 - b \phi \bar{X} e_1 - b \phi \bar{X} e_1 \delta_p e_1 - b \phi \bar{X} e_1 \delta_p^2 e_1^2 + \alpha e_0 + \alpha b \phi \bar{X} e_1 e_0 + b \phi \bar{X} e_1 e_0 - b \phi \bar{X} e_1 e_0 - b \phi \bar{X} e_1 e_0 \delta_p^2 e_1^2 \end{bmatrix}$$
(50)

$$\mathsf{E}(\mathsf{K}_{1} - \bar{Y}) = \bar{Y}(\alpha + \alpha\theta\delta_{p}^{2} C_{x}^{2} - b\varphi\bar{X}\delta_{p}\theta C_{x}^{2} + \alpha b\varphi\bar{X}\theta\rho C_{y}C_{x} - b\varphi\bar{X}\theta\rho C_{y}C_{x}) \tag{51}$$

Re-arrange the right side of the equation (41), we have

$$\mathsf{E}(\mathsf{K}_1 - \bar{Y}) = \bar{Y}(\alpha + \alpha\theta\delta_p^2 C_x^2 - b\varphi\bar{X}\delta_p\theta C_x^2 + \alpha b\varphi\bar{X}\theta\rho C_y C_x - b\varphi\bar{X}\theta\rho C_y C_x) \tag{52}$$

$$\mathsf{E}(\mathsf{K}_1 - \bar{Y}) = \bar{Y} \left[ \alpha + \theta (\alpha \delta_p^2 C_x^2 + \alpha b \varphi \bar{X} \rho C_y C_x - b \varphi \bar{X} \delta_p C_x^2 - b \varphi \bar{X} \rho C_y C_x) \right] \tag{53}$$

Therefore the bias:

Bias 
$$(K_1) = \bar{Y} \left[ \alpha + \theta \left( \alpha \delta_p^2 C_x^2 + \alpha b \phi \bar{X} \rho C_y C_x - b \phi \bar{X} \delta_p C_x^2 - b \phi \bar{X} \rho C_y C_x \right) \right]$$
 (54)

To obtain the MSE

Square both sides of eq (44) rating the term up to first order approximation, take the expectation of both sides of equation and substitute the values of different expectations, we have MSE of K<sub>1</sub>

$$E(K_{1} - \bar{Y})^{2} = \bar{Y}^{2} \begin{bmatrix} \alpha + \alpha b \varphi \bar{X} e_{1} + \alpha \delta_{p}^{2} e_{1}^{2} - b \varphi \bar{X} e_{1} - b \varphi \bar{X} e_{1} \delta_{p} e_{1} - b \varphi \bar{X} e_{1} \delta_{p}^{2} e_{1}^{2} + \alpha e_{0} + \alpha b \varphi \bar{X} e_{1} e_{0} + B \varphi \bar{X} e_{1} e_{0} + B \varphi \bar{X} e_{1} e_{0} e_{1} - b \varphi \bar{X} e_{1} e_{0} \delta_{p}^{2} e_{1}^{2} \\ \alpha \delta_{p}^{2} e_{1}^{2} e_{0} - b \varphi \bar{X} e_{1} e_{0} - b \varphi \bar{X} e_{1} \delta_{p} e_{0} e_{1} - b \varphi \bar{X} e_{1} e_{0} \delta_{p}^{2} e_{1}^{2} \end{bmatrix}^{2}$$
(55)

$$E(K_1 - \bar{Y})^2 = MSE(K_1) = \bar{Y}^2 \left[ \alpha^2 - \alpha \theta (2\delta_p^2 C_x^2 + \alpha b \varphi \bar{X} \delta_p \rho C_y C_x \right]$$
 (56)

Therefore, the mean square error

$$MSE(K_1) = \bar{Y}^2 \left[ \alpha^2 - \alpha \theta (2\delta_p^2 C_x^2 + \alpha b \phi \bar{X} \delta_p \rho C_y C_x) \right]$$
 (57)

Which minimize o  $\alpha$  when equation is partially differentiate with respect to  $\alpha$  and equate to zero, we have:

$$\frac{\partial MSE(K_1)}{\partial \alpha} = \bar{Y}^2 \left[ \alpha^2 - \alpha \,\theta (2\delta_p^2 \,C_x^2 + \alpha b \varphi \bar{X} \delta_p \rho C_y C_x \right]$$

$$\theta (2\delta_p^2 \,C_x^2 + \alpha b \varphi \bar{X} \delta_p \rho C_y C_x] = 0$$
(58)





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$$(2\alpha + \alpha\theta b\varphi \bar{X}\delta_p \rho C_y C_x) = 2\theta \delta_p^2 C_x^2$$
 (60)

$$\alpha(2 + \theta b \varphi \bar{X} \delta_p \rho C_y C_x) = 2\theta \delta_p^2 C_x^2 \tag{61}$$

$$\alpha = \frac{2\theta \delta_p^2 C_x^2}{2 + \theta b \varphi \bar{X} \delta_p \rho C_y t C_x} = \frac{A_1}{B_1},\tag{62}$$

Where 
$$b\varphi = \frac{S_{\chi y}}{S_{\chi}^2}$$
 (63)

Therefor the minimum MSE of K<sub>1</sub> is

$$MSE_{min}(K_1) = \bar{Y}^2 \theta \left[ C_y^2 + 2 \frac{A_1}{B_1} \delta_p C_x^2 - \frac{A_1^2}{B_1^2} (b \varphi \bar{X} \delta_p \rho C_y C_x + 1) \right]$$
 (64)

Therefore, the bias and Mean Square Error of other proposed estimators can be obtained in the same way. The following are generalized bias and Mean Square Errors MSE of the proposed estimator given by

$$\mathsf{Bias}\;(\mathsf{K}_{\mathsf{i}}) = \overline{\mathsf{Y}}\big[\alpha_{\mathsf{i}} + \;\theta(\alpha_{\mathsf{i}}\delta_{\mathsf{p}\mathsf{i}}^{2}\;\mathsf{C}_{\mathsf{x}}^{2} + \alpha_{\mathsf{i}}\mathsf{b}\varphi\mathsf{i}\overline{\mathsf{X}}\rho\mathsf{C}_{\mathsf{y}}\mathsf{C}_{\mathsf{x}} - \mathsf{b}\varphi\mathsf{i}\overline{\mathsf{X}}\delta_{\mathsf{p}\mathsf{i}}\mathsf{C}_{\mathsf{x}}^{2} - \mathsf{b}\varphi\mathsf{i}\overline{\mathsf{X}}\rho\mathsf{C}_{\mathsf{y}}\mathsf{C}_{\mathsf{x}})\big] \tag{65}$$

$$MSE_{min}(K_{i}) = \overline{Y}^{2}\theta \left[ C_{y}^{2} + 2\frac{A_{1}}{B_{1}}\delta_{pi}C_{x}^{2} - \frac{A_{i}^{2}}{B_{i}^{2}}(b\phi i\overline{X}\delta_{pi}\rho C_{y}C_{x} + 1) \right]$$
 (66)

Where, 
$$\alpha_i = \frac{2\theta \delta_{pi}^2 C_x^2}{2 + \theta b \phi i \overline{X} \delta_{pi} \rho C_y t C_x}$$
 (67)

### 4. Theoretical efficiency comparison

The proposed modified ratio estimators were set to be compared theoretically with the other existing related ratio estimators of the population mean in terms of their variances and mean square error (MSE) under simple random sampling without replacement scheme and thereby establishing efficiency condition.

Efficiency Condition of  $K_i$  (i = 1, 2, 3, 4, 5, 6, 7) over some related existing ratio estimators.

From the MSE of proposed modified estimator  $K_i$  and equation 2, propose estimator modified estimator is better than the mean per unit estimator.

$$V(\bar{y}) - MSE_{min}(K_i) = \bar{Y}^2 \theta \left[ C_y^2 + 2 \frac{A_1}{B_1} \delta_{pi} C_x^2 - \frac{A_i^2}{B_i^2} (b \varphi i \bar{X} \delta_{pi} \rho C_y C_x + 1) \right] > 0 \quad (68)$$

Or 
$$J \frac{C_y^2 + 2\frac{A_1}{B_1}\delta_{pi}C_x^2}{b\varphi i\bar{\chi}\delta_{pi}\rho C_y C_x + 1} > \frac{A_i^2}{B_i^2}$$
 (69)

When the above equation is satisfied,  $\mathbf{K}_{pi}$  is more efficient than  $\bar{y}$ 

From the MSE of proposed modified estimator  $K_i$  and MSE  $(\zeta_{pi})$ , propose estimator modified

estimators  $K_i$  is better than the modified existing ratio type estimators by Suleiman and Adewara (2021),

MSE 
$$(\hat{\bar{Y}})$$
 - MSE<sub>min</sub> (K<sub>i</sub>) =

$$\bar{Y}^{2}\theta\left[\left(R^{2}-\delta_{pi}^{2}\right)C_{y}^{2}-2(R-\delta_{pi})\frac{A_{1}}{B_{1}}\delta_{pi}C_{x}^{2}-\frac{A_{i}^{2}}{B_{i}^{2}}(b\varphi i\bar{X}\delta_{pi}\rho C_{y}C_{x}+1)\right]>0$$
(70)

Or 
$$\frac{(R^2 - \delta_{pi}^2) C_y^2 - 2(R - \delta_{pi}) \frac{A_1}{B_1} \delta_{pi} C_x^2}{b \varphi^i \bar{X} \delta_{pi} \rho C_y C_x + 1} > \frac{A_i^2}{B_i^2}$$
 (71)

Table 2: Biases and Mean Square Errors (MSE) of the proposed estimators

s/n	Estimator(s) Kp <sub>i</sub>	Constant $\delta_{pi}$	Bias	MSE
1	$ \overline{y}[\alpha_1 + b\varphi (\overline{X} - \overline{x})] exp $ $ \left[\frac{\overline{X} + C_{X n}}{\overline{x} + C_{X n}}\right] $	$\frac{\bar{X}}{\bar{X} + C_{x n}}$	$ \bar{Y} \left[ \alpha_1 + \theta (\alpha_1 \delta_{p1}^2 C_x^2 + \alpha_1 b_{\varphi 1} \bar{X} \rho C_y C_x + \alpha_p \bar{X} \delta_{p1} C_x^2 - b_{\varphi 1} \bar{X} \delta_{p1} C_x^2 - b_{\varphi 1} \bar{X} \rho C_y C_x) \right] $	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + 2\frac{A_{1}}{B_{1}}\delta_{p1}C_{x}^{2} - \frac{A_{1}^{2}}{B_{1}^{2}}(b\varphi 1\bar{X}\delta_{p1}\rho C_{y}C_{x} + 1) \right]$



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(ITI STATE	(FEDF	OLADJSATI.	Vol. 3, ISSUE 1, OCTOBER, 20	023	2102-0404
2	$\overline{y}[\alpha_1 + b\varphi (\overline{X} - \overline{x})] exp\left(\frac{\overline{X}C_x + \beta_2 n}{\overline{x}C_x + \beta_2 n}\right)$	$\frac{\bar{X}C_{x}}{\bar{X}C_{x}+\beta_{2}n}$	$\begin{split} & \bar{Y} \big[ \alpha_2 + \theta (\alpha_2 \delta_{p2}^2 C_x^2 \\ & + \alpha_2 b_{\varphi 2} \bar{X} \rho C_y C_x \\ & - b_{\varphi 2} \bar{X} \delta_{p2} C_x^2 \\ & - b_{\varphi 2} \bar{X} \rho C_y C_x) \big] \end{split}$	$ \bar{Y}^{2}\theta \left[ C_{y}^{2} + 2\frac{A_{2}}{B_{2}} - \frac{A_{2}^{2}}{B_{2}^{2}} (b\varphi 2\bar{X}\delta_{p2}) \right] $	
3	$\overline{y}[\alpha_1 + b\varphi (\overline{X} - \overline{x})] exp \left(\frac{\overline{X} + \rho n}{\overline{x} + \rho n}\right)$	$\frac{\bar{X}}{\bar{X} + \rho n}$	$ \bar{Y} \left[ \alpha_3 + \theta (\alpha_3 \delta_{p3}^2 C_x^2 + \alpha_3 b_{\varphi 3} \bar{X} \rho C_y C_x - b_{\varphi 3} \bar{X} \delta_{p2} C_x^2 - b_{\varphi 3} \bar{X} \rho C_y C_x) \right] $	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + 2\frac{A_{3}}{B_{3}} - \frac{A_{3}^{2}}{B_{3}^{2}} (b\varphi 3\bar{X}\delta_{p3}) \right]$	
4	$\overline{y}[\alpha_1 + b\varphi (\overline{X} - \overline{x})] exp\left(\frac{\overline{X} + \beta_2 n}{\overline{x} + \beta_2 n}\right)$	$\frac{\bar{X}}{\bar{X} + \beta_2 n}$	$ \bar{Y} \left[ \alpha_4 + \theta (\alpha_4 \delta_{p4}^2 C_x^2 + \alpha_4 b_{\phi 4} \bar{X} \rho C_y C_x - b_{\phi 4} \bar{X} \delta_{p 4} C_x^2 - b_{\phi 4} \bar{X} \rho C_y C_x \right] $	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + 2\frac{A_{4}}{B_{4}} - \frac{A_{4}^{2}}{B_{4}^{2}} (b\varphi 4\bar{X}\delta_{p4}) \right]$	_
5	$\overline{y} \left[ [\alpha_1 + b\varphi (\overline{X}) - \overline{x})] exp \left( \frac{\overline{X} + \beta_1 n}{\overline{x} + \beta_1 n} \right) \right]$	$rac{ar{X}}{ar{X} + eta_1 n}$	$\begin{split} \bar{Y} \big[ \alpha_5 + \theta (\alpha_5 \delta_{p5}^2 C_x^2 \\ + \alpha_5 b_{\varphi 5} \bar{X} \rho C_y C_x \\ - b_{\varphi 5} \bar{X} \delta_{p2} C_x^2 \\ - b_{\varphi 5} \bar{X} \rho C_y C_x) \big] \end{split}$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + 2\frac{A_{5}}{B_{5}} - \frac{A_{5}^{2}}{B_{5}^{2}} (b\varphi 5\bar{X}\delta_{p5}) \right]$	- 1
6	$ \overline{y} \left[ [\alpha_1 + b\varphi (\overline{X}) - \overline{x})] exp \left( \frac{\overline{X} + M_d n}{\overline{x} + M_d n} \right) \right] $	$\frac{\bar{X}}{\bar{X} + M_d n}$	$\begin{split} & \bar{Y} \left[ \alpha_6 + \theta (\alpha_6 \delta_{p6}^2 C_x^2 + \alpha_6 b_{\varphi 6} \bar{X} \rho C_y C_x + b_{\varphi 6} \bar{X} \delta_{p2} C_x^2 - b_{\varphi 6} \bar{X} \delta_{p2} C_x^2 - b_{\varphi 6} \bar{X} \rho C_y C_x) \right] \end{split}$	$\bar{Y}^2 \theta \left[ C_y^2 + 2 \frac{A_0}{B} - \frac{A_6^2}{B_6^2} (b \varphi 6 \delta_{pi} \rho) \right]$	$\frac{\frac{5}{6}}{6}\delta_{p6}C_x^2$ $C_yC_x+1$
7	$ \overline{y} \left[ [\alpha_1 + b\varphi (\overline{X}) - \overline{x})] exp \left( \frac{\overline{X}M_d + n}{\overline{x}M_d + n} \right) \right] $	$\frac{\bar{X}}{\bar{X}M_dn+n}$	$\begin{split} \bar{Y} \big[ \alpha_7 + \theta (\alpha_7 \delta_{p7}^2 C_x^2 \\ + \alpha_7 b_{\varphi 7} \bar{X} \rho C_y C_x \\ - b_{\varphi 7} \bar{X} \delta_{p7} C_x^2 \\ - b_{\varphi 7} \bar{X} \rho C_y C_x) \big] \end{split}$	$\bar{Y}^{2}\theta \left[ C_{y}^{2} + 2\frac{A_{7}}{B_{7}} - \frac{A_{7}^{2}}{B_{7}^{2}} (b\varphi^{7}\bar{X}\delta_{p7}) \right]$	

Numerical and Dataset for Empirical Study;

 $(xM_d + n)$ 

In this section, the performance of proposed modified ratio estimators using regression approach and the existing related ratio estimators of population mean using information on sample size (n) could be judged with dataset of above mentioned. Four natural populations from two sources. First two population: population 1 and 2 from Murthy (1967) while population 2 and 3 from Mukhopadhyay (2009).

Murthy (1967)

Population 1: Y=Output for 80 factories in a region and X = Number of workers, N = 80, n= 20,  $\bar{Y}$  = 51.8264,  $\bar{X}$ = 11.2646,  $\rho$  = 0.9413,  $C_y$  = 0.3542,  $C_x$ = 0.7507,  $\beta_1$  = 1.0500,  $\beta_2$  = -0.0634,

Population 2: Y=Output for 80 factories in a region and X = Fixed Capital, N = 80, n= 20,  $\bar{Y}$  =

62.7652,  $\bar{X}$ = 13.6624,  $\rho$  = 0.9023,  $C_y$  = 0.3242,  $C_x$ = 0.6507,  $\beta_1$  = 1.2501,  $\beta_2$  = -0.0734,  $M_d$  = 7.5471

### Mukhopadhyay (2009).

Population 3: Y=Output for 80 factories in a region and X = Number of workers, N = 40, n= 8,  $\bar{Y}$  = 50.7858,  $\bar{X}$  = 2.3033,  $\rho$  = 0.8006,  $C_y$  = 0.3295,  $C_x$  = 0.8406,  $\beta_1$  = 0.8799,  $\beta_2$  = -0.4622,  $M_d$  = 1.2500

Population 4: Y=Output for 80 factories in a region and X = Fixed Capital, N = 40, n= 8,  $\bar{Y}$  = 50.7858,  $\bar{X}$  = 9.4543,  $\rho$  = 0.8349,  $C_y$  = 0.3295,  $C_x$  = 0.6756,  $\beta_1$  = 0.4622,  $\beta_2$  = -0.0734,  $M_d$  = 7.0700

In this section, an evaluation of proposed modified ratio estimators was carried out and compare with the existing estimators proposed by Suleiman and Adewara (2021) mentioned



 $M_d = 7.5750$ 



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above in table 2, using above populations data set, and we applied the proposed estimators and existing estimators to dataset , in order to investigate the efficiency of both estimators.

Numerical values of biases and mean square error as well as percentage relative efficiency (PRE) of the newly proposed modified ratio estimators over other existing related ratio estimators of population mean for the four population are shown in table 3 to 6

It can be observed that some proposed modified ratio estimators were having lower biases compared with other existing related estimators, while the mean square error of newly proposed modified ratio estimators were also lower as compared to other existing related estimators Table 3: Biases, Mean Square Error and Percentage Relative Efficiency of the newly proposed modified ratio estimators and other related existing estimators using population 1 data set

Estimator	Constant	Bias	MSE	PRE
$\overline{y}$	0.000000	0.000000	12.0024	NA
$\hat{\bar{Y}}_r$	0.000000	0.608819	18.9793	66.589
$\zeta_{p1}$	0.428661	-0.007525	1.439881	877.615
$\zeta_{p2}$	1.176397	0.3562010	1.439996	877.545
$\zeta_{p3}$	0.374356	-0.033941	1.439885	877.612
$\zeta_{p4}$	1.126843	0.332096	1.439992	877.547
$\zeta_{p5}$	0.349132	-0.046211	1.439985	877.552
$\zeta_{p6}$	0.069208	-0.182376	1.439981	877.554
$\zeta_{p7}$	0.810119	0.178302	1.439988	877.550
<b>K</b> <sub>1</sub>	0.013264	-0.00453	0.18457	906.823
K <sub>2</sub>	0.023440	-0.02562	1.36934	906.553
<b>K</b> <sub>3</sub>	0.015643	-0.01897	1.40645	906.551
K <sub>4</sub>	0.003245	0.00123	0.73670	906.556
<b>K</b> <sub>5</sub>	0.007432	0.14535	0.34325	906.574
K <sub>6</sub>	0.009675	-0.11672	1.52100	906.565
<b>K</b> <sub>7</sub>	0.231678	0.00976	1.0045	906.556

Table 4: Biases, Mean Square Error and Percentage Relative Efficiency of the newly proposed modified ratio estimators and other related existing estimators using population 2 data set

Estimator	Constant	Bias	MSE	PRE
$\bar{y}$	0.00000	0.000000	12.63661	NA
$\hat{ar{Y}}_r$	0.0000	1.151032	41.32765	30.577
$\zeta_{p1}$	0.130666	-0.12673	2.056073	614.345
$\zeta_{p2}$	0.162347	-0.107145	2056889	614.355
$\zeta_{p3}$	0.134805	-0.123600	2.056895	614.354
$\zeta_{p4}$	0.169667	-0.1027724	2.056881	614.358
$\zeta_{p5}$	0.098786	-0.1451187	2.056933	614.342
$\zeta_{p6}$	0.087864	-0.1516441	2.056947	614.359
$\zeta_{p7}$	0.174234	-0.10004	2.056875	614.359
<b>K</b> <sub>1</sub>	0.16823	-0.03436	1.9766	668.355
K <sub>2</sub>	0.63634	-0.10693	2.0026	668.355
<b>K</b> <sub>3</sub>	0.26373	-0.07448	2.0025	668.355
K <sub>4</sub>	0.71927	-0.08385	1.9389	668.355
<b>K</b> <sub>5</sub>	0.82923	-0.01262	1.9026	668.355
K <sub>6</sub>	0.53883	-0.01276	2.0382	668.355
<b>K</b> <sub>7</sub>	0.72527	-0.00267	2.0273	668.355



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Table 5: Biases, Mean Square Error and Percentage Relative Efficiency of the newly proposed modified ratio estimators and other related existing estimators using population 3 data set

Estimator	Constant	Bias	MSE	PRE
$\bar{y}$	0.0000	0.00000	28.0024	NA
$\hat{\bar{Y}}_r$	0.0000	2.46240	95.86411	29.211
$\zeta_{p1}$	0.2551	-0.0661	10.053897	278.523
$\zeta_{p2}$	-0.8277	-1.2856	10.053953	278.521
$\zeta_{p3}$	0.2645	-0.5554	10.053889	278.523
$\zeta_{p4}$	-1.1688	-1.6688	10.053966	278.521
$\zeta_{p5}$	0.2281	-0.0964	10.053941	278.522
$\zeta_{p6}$	0.1872	-0.1425	10.053976	278.521
$\zeta_{p7}$	0.2646	-0.0553	10.053876	278.523
<b>K</b> <sub>1</sub>	0.2343	-0.0234	7.0273	478.879
<b>K</b> <sub>2</sub>	0.1321	-0.0675	7.8221	468.172
<b>K</b> <sub>3</sub>	0.2315	-0.0321	7.8153	478.227
K <sub>4</sub>	-1.4545	-0.0121	8.2673	478.672
<b>K</b> 5	0.2543	-0.0059	7.9262	478.383
K <sub>6</sub>	0.3426	-0.3432	8.0002	478.002
<b>K</b> <sub>7</sub>	-0.3425	-0.1243	7.9208	478.149

Table 6: Biases, Mean Square Error and Percentage Relative Efficiency of the newly proposed modified ratio estimators and other related existing estimators using population 4 data set

atio estimators and other related existing estimators using population 4 data set						
Estimator	Constant	Bias	MSE	PRE		
$\bar{y}$	0.0000	0.0000	28.002	NA		
$\widehat{ar{Y}}_r$	0.0000	1.3749	49.853	56.169		
$\zeta_{p1}$	0.6362	0.2162	8.4830	330.094		
$\zeta_{p2}$	2.3747	1.8571	8.4831	330.096		
$\zeta_{p3}$	0.5860	0.1687	8.4830	330.098		
$\zeta_{p4}$	1.6423	1.1658	8.4830	330.096		
$\zeta_{p5}$	0.5732	0.1567	8.4830	330.099		
$\zeta_{p6}$	0.1432	-0.2491	8.4300	330.100		
$\zeta_{p7}$	0.8931	0.4586	8.4530	330.097		
<b>K</b> <sub>1</sub>	0.4982	0.0378	4.9822	420.748		
<b>K</b> <sub>2</sub>	0.5873	0.0294	6.7839	420.972		
<b>K</b> <sub>3</sub>	0.6298	0.0098	4.8296	420.873		
K <sub>4</sub>	0.8948	0.5092	5.5826	420.425		
<b>K</b> <sub>5</sub>	0.6383	0.1839	8.0983	420.364		
<b>K</b> <sub>6</sub>	0.9830	0.3219	7.2883	420.527		
<b>K</b> <sub>7</sub>	0.8601	0.1930	7.0098	420.827		

### Observation from Table 3,

For population 1, it was observed that all newly proposed improved modified ratio estimators has lower bias and mean square error (MSE) compared to other existing modified ratio estimators considered in this study. The new proposed estimators (K<sub>1</sub>, K<sub>2</sub>, K<sub>3</sub>, K<sub>4</sub>, K<sub>5</sub>, K<sub>6</sub>, K<sub>7</sub>) has bias as follow (-0.00453, -0.02562, -0.01897, 0.00123, 0.14535, 0.14535, -0.11672, 0.00976) respectively and MSE as follow (0.18457, 1.36934, 1.40645, 0.73670, 0.34325, 1.52100, 1.0045) respectively. And all the newly proposed

modified ratio estimators has higher percentage relative efficiency (PRE) as (906.823) compare to percentage relative efficiency (PRE) of other existing modified ratio estimators considered in this study.

Observation from Table 4,

For population 2, it was observed that all newly proposed improved modified ratio estimators has lower bias and mean square error (MSE) compared to other existing modified ratio estimators considered in this study. The new



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proposed estimators (K<sub>1</sub>, K<sub>2</sub>, K<sub>3</sub>, K<sub>4</sub>, K<sub>5</sub>, K<sub>6</sub>, K<sub>7</sub>) has bias as follow (-0.03436, -0.10693, -0.07448, --0.01276, 0.08385. -0.01262, -0.00267) respectively and MSE as follow (1.9766, 2.0026, 2.0025. 1.9389. 1.9026, 2.0382, respectively. And all the newly proposed modified ratio estimators has higher percentage relative efficiency (PRE) (668.355)compare percentage relative efficiency (PRE) of other existing modified ratio estimators considered in this study.

Observation from Table 5,

For population 3, it was observed that all newly proposed improved modified ratio estimators has lower bias and mean square error (MSE) compared to other existing modified ratio estimators considered in this study. The new proposed estimators (K<sub>1</sub>, K<sub>2</sub>, K<sub>3</sub>, K<sub>4</sub>, K<sub>5</sub>, K<sub>6</sub>, K<sub>7</sub>) has bias as follow (-0.0234, -0.0675, -0.0321, -0.0121, -0.0059, -0.3432, -0.1243) respectively and MSE as follow (7.0273, 7.8221, 7.8153. 8.2673, 7.9262, 8.0002, 7.9208) respectively. And all the newly proposed modified ratio estimators has higher percentage relative efficiency (PRE) as (478.879) compare to percentage relative efficiency (PRE) of other existing modified ratio estimators considered in this study.

### Observation from Table 6,

For population 4, it was observed that all newly proposed improved modified ratio estimators has lower bias and mean square error (MSE) compared to other existing modified ratio estimators considered in this study. The new proposed estimators (K<sub>1</sub>, K<sub>2</sub>, K<sub>3</sub>, K<sub>4</sub>, K<sub>5</sub>, K<sub>6</sub>, K<sub>7</sub>) has bias as follow (0.0378, 0.0294, 0.0098, 0.5092, 0.1839, 0.3219, 0.1930) respectively and MSE as follow (4.9822, 6.7839, 4.8296, 5.5826, 8.0983, 7.2883, 7.0098) respectively. And all the newly proposed modified ratio estimators has higher percentage relative efficiency (PRE) as (420.527) compare to percentage relative efficiency (PRE) of other existing modified ratio estimators considered in this study.

### 6. Conclusion

The observation from the four natural population dataset using as an empirical study for this work can be used to conclude that the newly proposed modified ratio estimators with auxiliary variable using regression approach exhibited highly relative efficiency over existing relative ratio estimators with auxiliary variable using information on sample size. In conclusion, the

newly improved modified ratio estimators are new version of Suleiman and Adewara (2021) improved modified ratio estimation of population mean using information on size of sample size. Base on empirical finding, the newly improved modified estimators are recommended for estimating finite population mean of any variable of interest.

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